

# Sunspot equilibria with persistent unemployment fluctuations\*

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Very preliminary version

## Abstract

We provide a business cycle model able to replicate the large amount of persistence in output and unemployment fluctuations found in the data. These variations in the unemployment rate are the result of self-fulfilling changes in expectations about future inflation in the wage bargaining process between workers of firms.

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# 1 Introduction

Persistent unemployment affecting Europe since the mid 1970s remains one of the most discussed subjects in economics. Several alternative explanatory theories have been proposed. According to the structuralist approach the observed persistent increase in unemployment is explained by permanent shocks that raised the natural rate of unemployment.<sup>1</sup> In contrast, Blanchard and Summers (1986, 1987) introduced the hysteresis hypothesis, where unemployment persistence is explained by the strong dependence of unemployment on the history of transitory shocks. This can be due either to the existence of multiple equilibria (unit roots)<sup>2</sup>, or to the existence of an extremely slow speed of convergence to a unique equilibrium (quasi unit roots).<sup>3</sup> In this paper we simulate a model able to reproduce the observed persistence in unemployment that uses this last mechanism.

Recent studies, that have also used the second version of hysteresis to explain unemployment fluctuations, have introduced non competitive features in the labour market in an otherwise standard Real Business Cycles (RBC) model and considered serially correlated exogenous shocks on fundamentals, in particular technological shocks.<sup>4</sup> These works were able to match successfully some key stylized facts concerning the labour market, while remaining as successful as previous standard RBC models in explaining the more standard dimensions of the business cycle. However they relied on exogenous persistent technological shocks to account for short-run unemployment fluctuations and persistence. This may be hard to reconcile with the strong heterogeneity in unemployment fluctuations that many European countries have experienced (see Figure 1).<sup>5</sup> Furthermore, unemployment persistence is for a large part the result of the exogenous stochastic process rather than a pure internal propagation mechanism.

In this paper we offer an alternative explanation to unemployment persistence that does not rely on exogenous shocks. We rather rely on Keynes' original idea that changes in expectations or "animal spirits" may generate waves of optimism and pessimism that lead to aggregate fluctuations in out-

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<sup>1</sup>See Phelps (1994), Phelps and Zoega (1998).

<sup>2</sup>See for example Rocheteau (1999), Raurich et al (2006) for the first line of research.

<sup>3</sup>Empirical studies

<sup>4</sup>See among others Mertz (1995), Andolfato (1996), den Haan *et al.*(2000), Maffezzoli (2001), Chéron and Langot (2004).

<sup>5</sup>To be more precisely documented empirically.

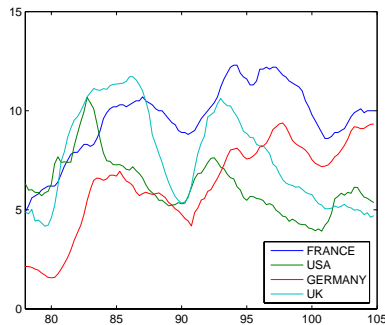


Figure 1: Unemployment rates in France, Germany, UK and USA, 1978:1 - 2004:4 (source: OEDC main indicators)

put and unemployment. This idea was revived in the eighties by the leading contributions of Azariadis (1981), Cass and Shell (1983), Azariadis and Guesnerie (1986), Grandmont (1985) and Woodford (1986), showing that, in the presence of indeterminacy of equilibria and/or bifurcations, sunspots shocks are able to generate endogenous cycles in economic variables. More recent research has also been studying how various market imperfections may favor the occurrence of indeterminacy and bifurcations leading to the existence of cycles driven by self fulfilling volatile expectations.<sup>6</sup>

Accordingly in this paper we consider an estimated version of the model we developed in Dufourt *et al.* (2005, forthcoming), where the Woodford (1986) framework of segmented asset markets with financially constrained workers is extended to account for several labor market imperfections, with unemployment emerging as an equilibrium result. In particular, we consider that wages and employment are bargained between unions and firms and that the government provides an imperfect insurance scheme ensuring a fixed minimum real income to the unemployed, financed by taxing employed workers. Since unions are able to set wages above the unemployment compensation, being unemployed is really costly in terms of welfare.

In accordance with our results previously obtained, this model is shown to be indeterminate for a large range of parameters values, including the empirically credible ones. In addition, Flip and Hopf bifurcations also occur within

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<sup>6</sup>Insert references.

this empirically credible range. This implies that stochastic fluctuations due to self-fulfilling changes in expectations are likely to emerge, leading to endogenous unemployment fluctuations that may have the appropriate degree of persistence. To verify this conjecture, we simulate and estimate a stochastic version of this model submitted to transitory (white noise) sunspot shocks, using a standard simulated method of moments and Minimum Distance criterion.

We find that the model is able to replicate the high persistence in unemployment fluctuations observed in many real economies, as measured by the autocorrelation function of the unemployment rate. It also replicates the autocorrelation function of output growth, a feature which is hardly accounted for by standard RBC models or models with endogenous fluctuations, as discussed in Cogley and Nason (1995) and Schmitt-Grohé (2000). This success occurs while, by definition, sunspot shocks are restricted to be i.i.d. stochastic processes. As we discuss, the main explanation for this success is that the estimated values for the unobserved structural parameters imply that the economy is located near to the point where a Hopf bifurcation occurs in the parameters space. At this point, the economy displays strong endogenous persistence.

The rest of the paper is organized as follows. Section 2 presents the model, which is essentially a simplified version of the more general framework studied in Dufourt *et al.* (2005, forthcoming). In section 3 we obtain the equilibrium and discuss the local dynamic properties of the model. Section 4 presents the estimation procedure and discusses the results, comparing them with the related literature. Finally section 5 concludes.

## 2 The Model

We consider a one sector heterogenous agents economy with segmented asset markets and costly unemployment, as developed in Dufourt *et al.* (forthcoming). The basis of this model is the finance constrained economy considered in Grandmont *et al.* (1998) and first proposed by Woodford (1986). In that economy all markets are perfectly competitive and there are two assets - money and productive capital, and two types of households, "workers" and "capitalists". Capitalists do not work and discount the future less than workers. The latter face a borrowing constraint implying that they cannot borrow against current and future income to finance current consumption.

Dufourt et al. (forthcoming) obtain costly equilibrium unemployment in this framework introducing an (imperfect) insurance scheme provided by the government in a economy where, due to union power, wages are set above the reservation wage. Below we provide a brief description of this model, where for simplicity we consider a CES parameterized version of the production function.

Capitalists are identical and maximize  $E \sum_{t=0}^{\infty} \beta^t \text{Log} c_t^c$ , where  $0 < \beta < 1$  is the discount factor, and  $c_t^c$  is consumption in period  $t$ . They face the following budget constraint  $p_t c_t^c + p_t k_{t+1}^c + m_{t+1}^c = p_t R_t k_t^c + m_t^c$ , where  $p_t$  is the price of output,  $k_{t+1}^c$  and  $m_{t+1}^c$  are respectively the capital stock and money holdings at the outset of period  $t+1$ ,  $R_t = (r_t + 1 - \delta)$  is the real gross rate of return on capital,  $r_t$  is the real rental rate of capital, and  $0 \leq \delta \leq 1$  is the capital depreciation rate.

It can easily be verified that under the condition  $R_{t+1} > E_t \{p_t/p_{t+1}\}$ , capitalists only save in the form of capital and hold no money ( $m_{t+1}^c = 0$ ). Hence, from the solution of the capitalists' problem, we have (see Woodford (1986))

$$c_t^c = (1 - \beta) R_t k_t^c \quad (1)$$

$$k_{t+1}^c = \beta R_t k_t^c. \quad (2)$$

In addition to these optimality conditions, the following transversality condition must be verified:

$$\lim_{t \rightarrow \infty} E_0 \left\{ \beta^t \frac{1}{c_t^c} \left( k_{t+1}^c + \frac{m_{t+1}^c}{p_t} \right) \right\} = 0 \quad (3)$$

There is a continuum of workers. Workers preferences are represented by the following utility function  $E \sum_{t=0}^{\infty} \gamma^t u(c_t)$ , where  $u(\cdot)$  is a standard quasi-concave real-valued continuous function,  $c_t$  is consumption in period  $t$  and  $0 < \gamma < \beta$  is the discount factor of workers. Each period, a worker supplies inelastically one unit of labor and may be either employed (state  $e$ ) receiving in cash, at the beginning of next period, a nominal wage  $w_t$ , or unemployed (state  $u$ ). If unemployed in  $t$  a worker receives from the government, at the beginning of period  $t+1$ , a constant real unemployment benefit  $b$ , financed by collecting, in period  $t+1$ , a given real amount from each worker employed at  $t$ . The budget constraint of a worker who was in state  $i \in \{e, u\}$  in period  $t-1$  can be written as  $m_{t+1}^i + p_t k_{t+1}^i = m_t + p_t y_t^i + p_t R_t k_t - p_t c_t^i$ , where  $m_t^w$  denotes money held at the beginning of period  $t$ , and where  $y_t^i \in \{w_{t-1} - p_t \tau_{t-1}, p_t b\}$ .

Additionally they face the borrowing constraint  $m_{t+1}^i \geq 0$ , and  $k_{t+1}^i \geq 0$  for all  $t$ .<sup>7</sup> Denoting by  $\lambda_t^i$ ,  $v_t^i$  and  $\eta_t^i$  the Lagrange multipliers associated respectively with these three constraints, the first order conditions for the workers' problem with a positive level of consumption are given by:

$$u'(c_t^i) = p_t \lambda_t^i \quad (4)$$

$$\lambda_t^i - v_t^i = \gamma E_t \{l_t \lambda_{t+1}^e + (1-l_t) \lambda_{t+1}^u\} \quad (5)$$

$$p_t \lambda_t^i - \eta_t^i = \gamma E_t \{p_{t+1} R_{t+1} [l_t \lambda_{t+1}^e + (1-l_t) \lambda_{t+1}^u]\} \quad (6)$$

where  $l_t$  denotes the employment rate in period  $t$ .

As in Dufourt et al. (forthcoming), it is easy to show from this system that if workers are significantly more impatient than capitalists (in a sense made more precise below), there exists equilibria along which both types of workers (employed and unemployed) rationally choose to have no saving and to spend all their available income for current consumption, i.e., for  $i = \{e, u\}$

$$c_t^i = \frac{y_t^i}{p_t} \quad (7)$$

with  $m_{t+1}^i = 0$  and  $k_{t+1}^i = 0$ .

It is easy to verify that such equilibria occur if sequences of revenues and probability distributions over employment and unemployment satisfy the following two conditions:

$$u'(c_t^i) > \gamma E_t \left\{ \frac{p_t}{p_{t+1}} [l_t u'(c_{t+1}^e) + (1-l_t) u'(c_{t+1}^u)] \right\} \quad (8)$$

$$u'(c_t^i) > \gamma E_t \{R_{t+1} [l_t u'(c_{t+1}^e) + (1-l_t) u'(c_{t+1}^u)]\} \quad (9)$$

Note that, from (2), the steady state interest factor is  $R = 1/\beta > 1$ , implying that the inequality  $R_{t+1} > E_t \{p_t/p_{t+1}\}$  holds in a neighborhood of this steady-state and that condition (9) is more restrictive than (8). Due to the concavity of  $u$ , and provided that  $y^e \geq y^u$  (a condition which, as we shall see, is implied by the wage bargaining process), condition (9) is in particular verified at the steady state and in its neighborhood iff

$$\gamma/\beta < u'(y^e) / (l u'(y^e) + (1-l) u'(y^u)). \quad (10)$$

implying that workers are sufficiently more impatient than capitalists.

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<sup>7</sup>For simplicity of notation, we dropped the superscrit  $w$ .

Each period, wages and employment are bargained between unions and firms. All workers are unionized and unions are firm-specific, i.e., there is one union per firm and each union represents the same mass of workers, normalized to one.<sup>8</sup> Unions wish to maximize the sum of discounted consumptions of their representative member, and firms wish to maximize the expected value of discounted profits,  $\Pi_t$ . Firms operate under a CES production function with constant returns to scale. Accordingly, we have that

$$F(k_t, l_t) = Al_t f(x_t) = Al_t \left[ \varphi x_t^{\frac{\sigma-1}{\sigma}} + (1 - \varphi) \right]^{\frac{\sigma}{\sigma-1}} \quad ; \quad 0 < \varphi < 1 \quad (11)$$

where  $x_t \equiv k_t/l_t$ ,  $k_t$  and  $l_t$  represent respectively capital and labour employed in each firm<sup>9</sup>,  $\sigma > 0$  is the constant factor elasticity of substitution and  $A > 0$  is a scale parameter. Each period  $t$  events follow the following sequence. First, firms pay in cash last period wages using previous money holdings. Then firms rent capital,  $k_t$ , at a given nominal rental rate  $p_t r_t$ . Next, wages,  $w_t$ , and employment,  $l_t$ , are negotiated between unions and firms. Finally, firms decide the amount of money holdings, given that it must be sufficient to at least cover the payment of wages,  $m_{t+1}^f \geq w_t l_t$ . Finally, production takes place. In order to ensure time consistency, the problem of the firm must be solved backwards, starting with the money holdings decision. As formally shown in Dufourt et al (forthcoming), it is straightforward to verify that the firms cash constraint is always binding, i.e.  $m_{t+1}^f = w_t l_t$ . We proceed now with the wage-employment bargain and then with capital decisions.

We model the bargaining process between each union and the respective firm using the standard generalized Nash bargaining solution. Let  $0 < \alpha \leq 1$  be the firm's power in the bargain, then the solution of the Nash maximization problem, assuming that  $l_t < 1$  is given by:<sup>10</sup>

$$(b + \tau_t) E_t \pi_{t+1} = A \left[ \varphi x_t^{\frac{\sigma-1}{\sigma}} + (1 - \varphi) \right]^{\frac{1}{\sigma-1}} (1 - \varphi) \quad (12)$$

$$\frac{w_t}{p_t} = A \left[ \varphi x_t^{\frac{\sigma-1}{\sigma}} + (1 - \varphi) \right]^{\frac{1}{\sigma-1}} \left[ (1 - \alpha) \varphi x_t^{\frac{\sigma-1}{\sigma}} + (1 - \varphi) \right] \quad (13)$$

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<sup>8</sup>Workers are matched exogenously and uniformly with unions and cannot move between firms or unions

<sup>9</sup>As we have normalized the mass of workers per firm to one,  $l$  represents both the employment level in each firm and the employment rate in the economy at a symmetric equilibrium.

<sup>10</sup>See Dufourt et al (forthcoming) for the derivation.

where  $\pi_{t+1} \equiv p_{t+1}/p_t$  is the inflation factor. From (12) we see that employment is determined by the equality between the firm's marginal productivity of labor,  $MPL = A \left[ \varphi x_t^{\frac{\sigma-1}{\sigma}} + (1-\varphi) \right]^{\frac{1}{\sigma-1}} (1-\varphi)$ , and the expected real reservation wage,  $(b + \tau_t)E_t\pi_{t+1}$ , for all  $\alpha$ . From (12) and (13) we see that, if  $\alpha < 1$ , unions are able to set wages above the reservation wage, i.e.,  $w_t/p_t > (b + \tau_t)E_t\pi_{t+1}$ , with a markup factor  $\left[ 1 + (1-\alpha)(\varphi/1-\varphi)x_t^{\frac{\sigma-1}{\sigma}} \right]$  which is increasing with union bargaining power. Hence, given the absence of perfect redistributive schemes, unemployed workers are clearly worse off. Another issue worth noting is that the level of employment is nonpredetermined, being influenced by expectations of future prices (or inflation). A change in expected future prices changes the reservation wage, and thereby the equilibrium level of employment.

The firm, anticipating the result of the bargaining process, chooses  $k_t > 0$  to maximize profits, which yields the following first order condition:

$$r_t = \varphi\alpha A \left[ (1-\varphi)x_t^{-\frac{(\sigma-1)}{\sigma}} + \varphi \right]^{\frac{1}{\sigma-1}}. \quad (14)$$

i.e.,  $r_t/\alpha$  equals the marginal productivity of capital.<sup>11</sup>

### 3 Equilibrium and dynamics

Since the government balances its budget, this real tax  $\tau_t$ , paid by each worker employed in period  $t$ , is determined endogenously by the balanced-budget condition

$$\tau_t = b(1 - l_t)/l_t \quad (15)$$

Assuming, as in Woodford (1986), a constant (per firm) amount of outside money in the economy,  $m$ , and given that only firms hold money, money market equilibrium in every period implies that  $m = w_t l_t = w_{t+1} l_{t+1}$ , so that  $p_{t+1}/p_t = (w_t l_t / p_t) / (w_{t+1} l_{t+1} / p_{t+1})$ . Using this last relation, equations (15), (12) and (13) we obtain equation (17) below. Considering an identical number of firms and capitalists, and using the definition of  $R$  and equations

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<sup>11</sup>Note that we recover the competitive outcome when unions have no power in the bargaining process ( $\alpha = 1$ ): i.e., the real rental cost of capital is identical to the marginal productivity of capital, and the real wage equals both the real reservation wage and the marginal productivity of labor.



(2) and (14), equilibrium in the capital services market implies equation (16) below. Accordingly we have:

**Definition 1** *A rational expectations intertemporal equilibrium is a sequence  $(k_t, l_t) \in \mathfrak{R}_{++}^2, t = 1, 2, \dots, \infty$  that solves the two-dimensional dynamic system, with  $x_t = k_t/l_t$*

$$k_{t+1} = \beta \left[ \varphi \alpha A_t \left[ (1 - \varphi) x_t^{-\frac{(\sigma-1)}{\sigma}} + \varphi \right]^{\frac{1}{\sigma-1}} + (1 - \delta) \right] k_t \quad (16)$$

$$b \left[ 1 + \frac{(1 - \alpha) \varphi x_t^{\frac{\sigma-1}{\sigma}}}{(1 - \varphi)} \right] = E_t \left\{ l_{t+1} A_{t+1} \left[ \varphi x_{t+1}^{\frac{\sigma-1}{\sigma}} + (1 - \varphi) \right]^{\frac{1}{\sigma-1}} \left[ (1 - \alpha) \varphi x_{t+1}^{\frac{\sigma-1}{\sigma}} + (1 - \varphi) \right] \right\} \quad (17)$$

Equation (16) represents the standard capital accumulation process, while equation (17) represents the equilibrium intertemporal arbitrage condition for workers.

To facilitate the discussion, it is useful to write this dynamic system under the following implicit form

$$E_t g(z_{t+1}, z_t) = 0 \quad (18)$$

where  $z_t \equiv (k_t, l_t)$ . It can be verified that, under non-restrictive conditions on parameters, this dynamic system has a unique steady state equilibrium  $\bar{z}$  defined by  $g(\bar{z}, \bar{z}) = 0$ , and that the Jacobian matrix of the map  $z \rightarrow g(z, \bar{z})$  evaluated at  $\bar{z}$  is invertible. By the Implicit Function Theorem, this system can therefore be solved in  $z_{t+1}$  in the neighborhood of  $\bar{z}$ , leading to a solution of the form

$$z_{t+1} = h(z_t, \epsilon_{t+1}) \quad (19)$$

where  $\epsilon_{t+1}$  is a vector of forecast errors.

### 3.1 Analysis of equilibria

From Definition 1, it is clear that  $k_t$  is a predetermined variable whose behavior is determined by past savings decisions of capitalists. However,  $l_t$  is a non-predetermined variable whose level is influenced by expectations. Hence, depending on the local stability properties of the steady-state, there is potentially the room for stationary stochastic equilibria driven by self-fulfilling

changes in expectations (sunspot shocks). We now briefly analyze when such situations can occur.

Consider first the case where the steady state is a saddle (locally determinate).<sup>12</sup> In this case, it is easily verified that expectations-driven equilibria can never arise. This is because, given the initial capital stock  $k_0$ , there is a unique locus defined by (19) which further satisfies the transversality condition (3) and is compatible with a long run convergence to the steady state. This implies that the forecast error  $\epsilon_{t+1}$  is necessarily zero in the absence of exogenous shocks on fundamentals.

The situation is completely different, however, when the steady state is a sink (or locally indeterminate). In this configuration, given the initial value of the capital stock  $k_0$ , there are now infinitely many equilibrium trajectories  $\{l_t, k_{t+1}\}_{t=0, \dots, \infty}$  compatible with (3) and converging to the steady state. Also, as proved by Azariadis and Guesnerie (1986), there are also infinitely many nondegenerate stochastic equilibria driven by self-fulfilling changes on expectations (sunspots equilibria), that stay arbitrarily close to the steady state. In terms of equation (19), this implies that the forecast error  $\epsilon_{t+1}$  may now act as an independent source of the business cycle, even in the absence of extrinsic uncertainty affecting fundamentals (see Benhabib and Farmer (1999) for further discussion on this point).

Finally, a last, but nonetheless interesting, type of equilibria is worth discussing. It can occur when the steady state is a source and the economy is located near the point where a supercritical Hopf bifurcation occurs in the parameters space.<sup>13</sup> In this case, as discussed in Grandmont et al. (1998), deterministic and (possibly) stationary stochastic equilibria generated by periodic or quasi-periodic orbits appear, lying around an invariant closed curve that surrounds the steady state in the state space. Thus, in this configuration, the economy may very well exhibit infinitely recurrent fluctuations in the unemployment rate *even in the absence of any kind of stochastic shocks* (on fundamentals or on expectations) - a form of "hysteresis" which

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<sup>12</sup>Explain

<sup>13</sup>A bifurcation occurs when the local stability properties of the system are drastically affected by a small change in parameters. Technically, it occurs when one (or several) eigenvalues of the characteristic polynomial of the system crosses the unit circle through this change of parameters. For example, a *flip* bifurcation occurs when one eigenvalue crosses -1. A *Hopf* bifurcation occurs when two complex conjugate eigenvalues have their modulus crossing 1.

is relatively new compared to the traditional literature.<sup>14</sup>

### 3.2 Dynamic configurations

In Dufourt *et al.* (forthcoming), a complete analytical characterization of the local stability properties of a (more general) version of this model has been undertaken in terms of relevant parameters. Figure 2, which is easily computed as a direct application of this theoretical analysis, reports in the  $(\sigma, \alpha)$  plane the bifurcation values for the elasticity of substitution  $\sigma$  as a function of the firms' bargaining power  $\alpha$ , given an empirically based calibration for the set of other parameters described below. Several interesting results can be drawn from the simple observation of this figure. First, the local dynamics of the model is indeterminate for a wide range of parameters values, including the empirically relevant ones. In particular, as proved in Dufourt *et al.* (forthcoming), the steady state is *always* indeterminate when the production function is Cobb-Douglas ( $\sigma = 1$ ). When the elasticity of substitution between capital and labor is different from one, both flip and Hopf bifurcations may occur. These bifurcations arise for empirically plausible values of  $\sigma$  (not far away from one) as soon as the unions' bargaining power is strong enough.<sup>15</sup>

In light of this analysis, it is clear that the two kinds of "sophisticated" dynamics resulting from the existence of sunspot equilibria with self-fulfilling changes in expectations, or from the existence of quasi periodic deterministic equilibria fluctuating along an invariant curve, are concrete possibilities, since they occur for plausible values of the structural parameters. For example, Figure 3 displays the closed invariant curve emerging in the model when the economy is located near the point where a Hopf bifurcation occurs. The Hopf bifurcation is found to be supercritical, implying that the invariant curve

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<sup>14</sup>To be more precisely explained

<sup>15</sup>In a recent paper, Grandmont (2007) introduced efficiency wages consideration rather than collective bargaining in an otherwise similar framework. His results show that the strong indeterminacy properties emphasized here are robust to changes in the wage setting process, making it clear that it is mostly the taxation/imperfect insurance scheme which is to be held responsible for indeterminacy. Interestingly, Grandmont (2007) also shows that increasing unemployment insurance may in some cases be welfare increasing along the deterministic stationary state, but that, in the same time, it increases the likelihood of indeterminacy (and, therefore, inefficient sunspot fluctuations). For a more detailed discussion of why indeterminacy easily occurs in this kind of models, see Dufourt et al. (forthcoming).

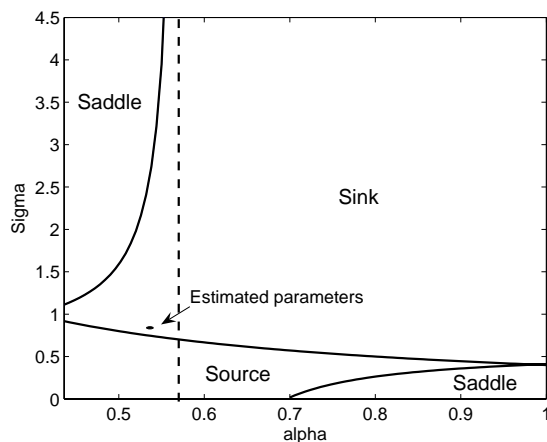


Figure 2: Local dynamics properties and bifurcations values in the space of parameters  $(\alpha, \sigma)$ , for given (calibrated) values of the set of parameters  $\gamma_1 = (\beta, \delta, b, \varphi)$ .

appears when the steady state is a source and that these curve is attracting.

In our view, both types of equilibria would have been worth studying. Unfortunately, the kind of strongly non-linear dynamics occurring around the invariant closed curve cannot always be easily analyzed by standard resolution techniques, including recent numerical ones.<sup>16</sup> For this reason, in order to study the empirical predictions of the model in terms of unemployment fluctuations and persistence, we have chosen to retain the more traditional approach pioneered by Benhabib and Farmer (1994) and Farmer and Guo (1994), consisting of generating fluctuations around the steady state due to self-fulfilling changes in expectations (sunspots), in an economy which is locally indeterminate.

## 4 Model evaluation

We now wish to investigate whether this model with self-fulfilling changes in expectations (sunspots shocks) can generate persistent, empirically con-

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<sup>16</sup>However, we were able to verify that the invariant closed curve appear near the bifurcation point when the steady-state is a source, implying that the bifurcation is supercritical.

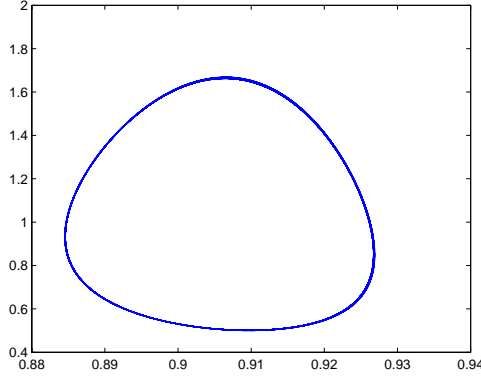


Figure 3: The invariant closed curve near the (supercritical) Hopf bifurcation.

sistent, fluctuations in the unemployment rate and output growth. In order to do so, an approximation of the solution to the dynamic system (16) and (17) is needed. Since we wish to consider the possibility that the economy be located near the points where the flip and Hopf bifurcations occur, it might be the case that the true dynamics of the model is too rich to be approximated sufficiently well by a standard linearization procedure. For this reason, we have followed instead the suggestion of Schmitt-Grohé and Uribe (2004) of approximating the solution using a second-order expansion of the policy function. This is likely to better capture the nonlinearities of the model, as we will show below.

When the steady state is a sink, sunspot equilibria driven by self-fulfilling changes in expectations exist, and a second order Taylor expansion of a solution satisfying (19) may be written as

$$\begin{aligned}\widehat{k}_{t+1} &\simeq \nabla_1 \widehat{z}_t + \frac{1}{2} \mathcal{Z}'_t H_1(\bar{z}) \widehat{z}_t \\ \widehat{l}_{t+1} &\simeq \nabla_2 \widehat{z}_t + \frac{1}{2} \mathcal{Z}'_t H_2(\bar{z}) \widehat{z}_t + \epsilon_{t+1}\end{aligned}$$

where  $\widehat{z}_t \equiv (\widehat{k}_t, \widehat{l}_t)$  is the vector of endogenous variables expressed in percentage deviations from the steady-state,  $\epsilon_{t+1}$  is a sunspot shock of bounded support with variance  $\gamma$ ,  $\nabla_i(\bar{z})$  is the gradient of the  $i^{th}$  component of  $h(\cdot)$ ,  $i = 1, 2$ , evaluated at the steady-state (or, equivalently, the  $i^{th}$  row of the

Jacobian matrix of  $h(\cdot)$  evaluated at  $\bar{z}$ ) and  $H_i$  is the Hessian matrix of  $h(\cdot)$  relative to variable  $i = 1, 2$  (or, equivalently, the jacobian matrix of  $\nabla_i(\cdot)$ ) evaluated at  $\bar{z}$ .

## 4.1 Calibration and estimation procedure

In order to simulate the model and evaluate its capacity to match empirical regularities, a sensible parameterization is needed. The model contains, besides the scale parameter  $A$ , six structural parameters:  $\beta$ ,  $\delta$ ,  $\sigma$ ,  $\alpha$ ,  $b$  and  $\varphi$ . Our general strategy is to partition these parameters into two groups: those for which there exists relatively common and rather uncontroversial estimates in the literature, or for which we can match balanced growth path values with observed averages ; and those for which such estimates are not available or are more controversial. The first set of parameters is calibrated, while the second set is chosen so as to minimize a measure of the distance between some preselected moments characterizing our data set and their model-implied counterparts.

The first set of parameters is  $\gamma_1 = (\beta, \delta, b, \varphi)$ . As we define the time period to be a quarter, we set  $\beta = 1.03^{-0.25}$ , which implies a steady state annualized real interest rate of 3 percent. We set  $\delta = 0.025$ , which implies an annual depreciation rate on capital of 10 percent. We calibrate the real amount of unemployment compensation  $b$  and the (unobserved) technological parameter  $\varphi$  so as to match the long-run labor share of output in France,  $s_L = 0.6$ , and the long-run level of unemployment over the period 1978:1 to 2004:4:  $u = 9.6\%$ .<sup>17</sup>

The second set of parameters includes the firms' bargaining power  $\alpha$  and the elasticity of substitution between capital and labor  $\sigma$ ,  $\gamma_2 = (\sigma, \alpha)$ . As these parameters are hardly observed or estimated, we follow Rotemberg and Woodford (1997), Christiano *et al.* (2005), and others, by estimating these parameters so as to match as closely as possible the preselected set of empirical moments using a Minimum Distance Estimation (MDE) procedure. To be more precise, let  $\Psi_T^e$  be a set of empirical moments characterizing our data set of length  $T$ , and let  $\Psi(\gamma_2)$  be the mapping from the (non calibrated) structural parameters to the corresponding theoretical set of moments. The Minimum Distance Estimator of  $\gamma_2$ , denoted  $\hat{\gamma}_2$ , is given by

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<sup>17</sup>Write note on the steady-state capital-labor ratio.

$$\hat{\gamma}_2 = \arg \min_{\gamma_2 \in \Gamma} (\Psi_T^e - \Psi(\gamma_2))' W (\Psi_T^e - \Psi(\gamma_2))$$

where  $W$  is a positive definite weighting matrix.

A problem that may arise in practice is that, given the relatively small number of observations in our data set ( $T = 108$ ), the model-generated sample equivalent of  $\Psi_T^e$  may be quite different from the theoretical one,  $\Psi(\gamma_2)$ . For this reason, we relied instead on a standard Method of Simulated Moments, where a short sample equivalent of  $\Psi(\gamma_2)$ , denoted  $\hat{\Psi}_T(\gamma_2)$ , was determined by repeatedly generating from the model artificial data sets of length  $T$  and then averaging the sample estimates. These repeated simulations were also used to compute an estimate  $\hat{\Sigma}$  of the variance-covariance matrix of  $\hat{\Psi}_T(\gamma_2)$ , which served as a basis for the confidence bounds below. Following Christiano et al. (2005), we chose as weighting matrix a diagonal matrix containing along the diagonal the inverse of the sample variances of  $\hat{\Psi}_T(\gamma_2)$ , i.e. the inverse of the diagonal elements of  $\hat{\Sigma}$ . With this choice, the vector or parameters  $\gamma_2$  is chosen so that the empirical moments  $\Psi_T^e$  lie as much as possible in these confidence bounds.

Finally, some discussion is required about the set of moments that we aimed to match. As the main issue of our paper is on unemployment persistence, and persistence in general, we have chosen to match the two statistical measures which emphasize the most strongly this dimension on the data. Namely, we have chosen to match the autocorrelation functions of the (HP-filtered) unemployment rate and of output growth of the French economy. In addition, a choice had to be made about the number of lags in the ACFs to consider. As the ACF function of output growth essentially vanishes after 6 lags, we chose as a benchmark to retain the first 6 lags of these autocorrelation functions. Results were not substantially altered, however, when we experimented with different numbers of lags.

## 4.2 Estimation results

The estimated vector of parameters was obtained based on the previously described minimization procedure using 600 simulations of data sets including  $T = 108$  observations. Table 1 reports the estimated values for  $\hat{\gamma}_2 = (\hat{\sigma}, \hat{\alpha})$ <sup>18</sup>, while Figure 4 reports the empirical autocorrelation function together with

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<sup>18</sup>Table 1 to be inserted. Compute standard deviations.

their theoretical (sample average) counterpart. Overall, the match appears very good, with the empirical and theoretical autocorrelation functions being very close from each other, and the two empirical autocorrelation functions lying entirely within the simulated 95% confidence bounds.<sup>19</sup>

$\hat{\sigma}$	$\hat{\alpha}$
0.741	0.454
(-)	(-)

Table 1 : Estimated parameters values

Hence, the simulated version of the model appears to be able to replicate the large amount of persistence in unemployment fluctuations and output growth which characterizes the French economy (and many similar countries). Note that these strongly persistent effects of shocks occur while, by definition, sunspot shocks are restricted to be *i.i.d* stochastic processes. There is therefore no doubt that this large amount of persistence is *endogenous* to the model, resulting entirely from internal propagation mechanisms and not from an exogenous source of persistence introduced through the stochastic driving processes. This is an important point, because Cogley and Nason (1995) and Rotemberg and Woodford (1996) strongly pointed out to the difficulties of many DSGE models to replicate the autocorrelation function of output growth without introducing an exogenous source of persistence<sup>20</sup>.

Another easy way to emphasize this strong endogenous persistence is to look at the estimated values for the set of parameters  $\gamma_2$ . Observe in this respect that the estimated vector  $\hat{\gamma}_2 = (\hat{\sigma}, \hat{\alpha})$  falls very close to the locus characterizing the Hopf bifurcation values for  $\sigma$  (see Figure 2). It is clear in this case that the model will display large endogenous persistence. In fact, in the context of our model, it is almost tautological to say that the model displays strong endogenous persistence or to emphasize that it falls close to a Hopf bifurcation in the parameters space. Indeed, persistence in a dynamic model can be analyzed by referring to the roots (eigenvalues)

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<sup>19</sup>To be more precise, 95% of the observed point estimates of the model-implied autocorrelation functions lie within the two dotted lines in figure 3.

<sup>20</sup>Of course, since the time of publication of these two papers, several extensions to the standard model have since been considered to try to improve these deficiencies of the original model. Relevant mechanisms include, without exhaustibility, factor hoarding (Burnside et al., 1996), etc... - include other references-. See the survey by King and Rebelo (1999) for more discussion on this issue.. As far as we know, however, very few papers can replicate these observations using only i.i.d. sunspot shock.



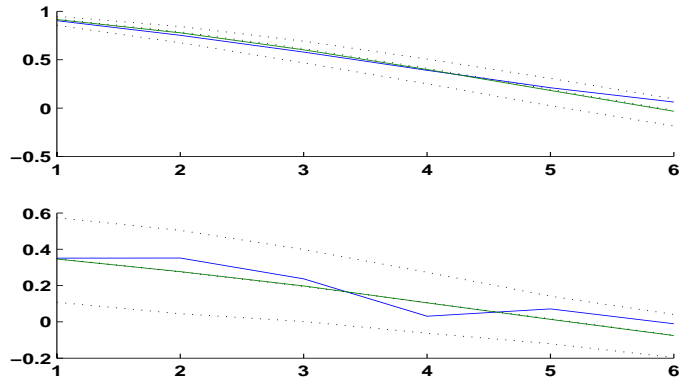


Figure 4: Empirical and model-implied autocorrelation functions for hp-filtered unemployment (upper graph) and output growth (lower graph). Dotted lines are confidence bounds.

of its characteristic polynomials. As our (reduced) dynamic model is two-dimensional, in a sink configuration, persistence requires that both roots have their modulus close to one. But this is precisely what is occurring when the model is in the neighborhood of a Hopf bifurcation, since in this case two complex conjugates eigenvalues of the dynamic system cross the unit circle under a small parameter change.

The important point to emphasize is that the Hopf bifurcation occurs for *realistic* values of all structural parameters. For example, the estimated value for the elasticity of substitution between capital and labor,  $\hat{\sigma} = .74$ , is very close to the value of  $\sigma = 0.7$  used in Pissarides (1997), Maffezzoli (2001), Cheron and Langot (2004), and others. Similarly, the estimated value for the firms bargaining power ( $\hat{\alpha} = 0.545$ ) is close to the standard value of 0.5 usually considered in the Labor Economics literature. It is also close to the value of 0.6 considered in the Real Business Cycle literature with wage bargaining (see e.g. Andolfatto (1996), Cheron and Langot, 2004).

This aspect is, we believe, one important contribution of our model with respect to the literature. In fact, in formal terms, the endogenous persistence resulting from our model occurs for similar reason to other papers in the literature. It results from the fact that the parameters values are such that the economy is close to the point where the Hopf bifurcation occurs.

However, in many papers, bifurcations and indeterminacy can only occur under rather controversial calibrations of parameters. This includes strong enough increasing returns to scale in production (see e.g. Benhabib and Farmer (1994), Farmer and Guo (1994), Barinci and Chéron (2001), Wen (1998), include others), distortive taxation (insert reference) or a high share of public spending in production (Schmitt-Grohe and Uribe (1997)).<sup>21</sup> In this model, by contrast, indeterminacy typically prevails under constant returns to scale and an *arbitrary* (positive) size of public redistribution.<sup>22</sup> Furthermore, provided that the unions' bargaining power is strong enough, the Hopf bifurcation arises for plausible values for the capital-labor elasticity of substitution .

Note also that the model offers an explanation to the high persistence of unemployment fluctuations which is relatively new in the literature. As mentioned in the introduction, early explanations for this feature have typically relied on *hysteresis* models with multiple equilibria, such as the "insiders-outsiders" types of models, in which the preferences of unions are implicitly assumed to exclude previously fired or unemployed workers.<sup>23</sup> In this tradition of models, persistent unemployment fluctuations occur because transitory shocks affect permanently the long run (or natural) level or the unemployment rate. While this type of explanations has received a great attention in the literature, the empirical evidence trying to assert it was at best mitigated. In fact, in many countries, different statistical tests applied to different economies often led to a rejection of a unit root in the unemployment series, suggesting a rather stable natural rate of unemployment (see e.g. Evans, 1989)

More recently, dynamic general equilibrium models in the RBC tradition have also attempted to account for the persistence in unemployment fluctuations without giving up the assumption of a unique (or stable) natural unemployment rate. Various frictions on the labor market have been considered. For example, Merz (1995), Andolfatto (1996), Gomes *et al.* (2001) considered frictions in the matching process between workers and firms. Maffezzoli

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<sup>21</sup>Wen (1998) is an interesting example of a model where the Hopf bifurcation arises for reasonable degrees of increasing returns to scale (higher than 10%). Benhabib and Wen (2004) have shown that this model could explain many features of the US business cycles. We discuss how our paper compares to this model in the next section.

<sup>22</sup>See Dufourt et al. (2005) for a more in depth discussion as well as an explanation for why indeterminacy occurs.

<sup>23</sup>See in particular Blanchard and Summers (1986).

(2001), Cheron and Langot (2004) introduced wage bargaining. In general, these papers showed that standard Dynamic Stochastic General Equilibrium (DSGE) models could reproduce the amount of unemployment persistence found in the data as long as persistent exogenous shocks (in particular technological innovations) were introduced as driving processes. They further showed that labor markets frictions were able to magnify the effects of these persistent exogenous shocks.<sup>24</sup>

Two important features differentiate our model to those of this literature. First, it is clear that persistent unemployment fluctuations can occur in our model *even in the absence of any shocks on fundamentals*. Rather, pessimistic or optimistic expectations of consumers and firms may help explain transitory but persistent fluctuations in the unemployment rate. In fact, business cycles practitioners often pay a lot of attention to the "confidence indices" of firms and consumers, such as those provided by the University of Michigan, because they know that these indices are reliable leading indicators of the business cycle. Our model is fully consistent with this view, as persistent fluctuations may be the result of autonomous changes in expectations. To our knowledge, this model is the first to account for persistent unemployment fluctuations resulting entirely from self-fulfilling changes in expectations.<sup>25</sup>

However, although we emphasize sunspots shocks as a potential source of unemployment fluctuations, it should be clear that such long lasting variations in the unemployment rate would result in this model from *any kind* of shocks (whether on fundamentals or on expectations) and *whatever the degree of persistence of these shocks*. Using white noise sunspot shocks is simply the most eloquent way to emphasize this dimension. We, of course, do not claim that sunspot shocks are the only source of fluctuations. But they may explain why similar countries with roughly similar economic conditions (as is the case for many European countries) have often experienced drastically different episodes of unemployment fluctuations.<sup>26</sup>

### 4.3 Other business cycles features

The ability of our model to account for the kind of strongly persistent fluctuations in the unemployment rate and output growth would be undermined

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<sup>24</sup>Describe more?

<sup>25</sup>Note also that, in contrast to these former models, the lack of a full insurance mechanism available to workers implies that being in unemployment is costly in terms of welfare.

<sup>26</sup>Insert examples.

if the model failed importantly in other dimensions of the business cycles. For this reason, we now turn to the evaluation of this model with respect to other standard features of the business cycle. This is an important step, because Schmitt-Grohé (2000) emphasized the difficulties of many DSGE models with sunspot-driven fluctuations to account for several stylized facts of actual economies, in particular (i) the inability of these models to account for the positive autocorrelation of output growth, (ii) their inability to reproduce the hump shaped response of output to transitory shocks, and (iii) the inability of most models to account for the procyclical response of consumption to transitory shocks, unless incredibly high increasing returns are assumed.

In a recent paper, Benhabib and Wen (2004) showed that many of these deficiencies could be alleviated by considering a version of Wen's (1998) model with variable utilization rate, moderate increasing returns to scale and exogenous (serially-correlated) variations in aggregate demand, resulting from preferences or government spending shocks. When their model is calibrated, as in our case, near the point where a bifurcation occurs, they show that it can display the correct amount of persistence in output growth, as measured by the different but related statistics emphasized by Cogley and Nason (1995) and Rotemberg and Woodford (1996). What is clear from their analysis is that these successes rely very much on the capacity of their model to generate a *hump-shaped* response of output to transitory/demand shocks. This is the case in their model, under the calibration considered, when indeterminacy is combined with exogenous, serially correlated, transitory shocks such as variations in preferences or government spending.

It is therefore important to relate our paper to Schmitt-Grohé (2000) and Benhabib and Wen (2004) in order to emphasize the similarities and differences.<sup>27</sup> The impulse response functions (IRFs) of the main variables to a sunspot shock obtained with the second-order approximation of the solution are displayed in Figure 3.<sup>28</sup> As expected, we can observe from Figure 3 that

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<sup>27</sup>It should be noted that we will mostly refer to the two stylized facts emphasized by Cogley and Nason (1995), regarding the autocorrelation of output growth and the hump shaped response of output to demand shocks, and not to the "forecastable movements" in the variables emphasized by Rotemberg and Woodford (1996), because the latter measures mostly make sense in the presence of persistent technological shocks affecting the long run level of output.

<sup>28</sup>For clarity of the figure, we have plotted the impulse response function of the employment rate, instead of that of the unemployment rate, because the latter series is by

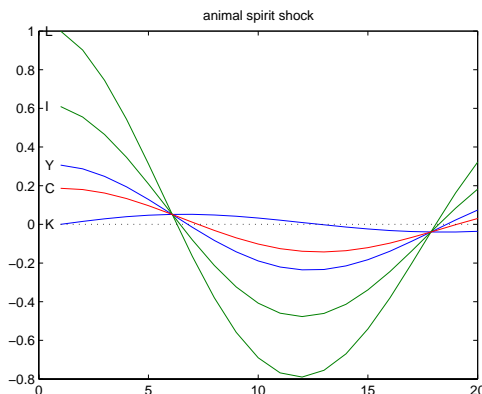


Figure 5: Impulse Response Functions with second-order approximation

sunspot shocks generate highly persistent periods of booms and recessions affecting simultaneously all the variables. In particular, in the aftermath of a positive sunspot shock, the employment rate, output, total consumption, investment and the capital stock all increase simultaneously for several periods, then decrease towards negative values for a few periods, and eventually revert back slowly to the initial steady state. Observe that deviations from the steady state are still significant even after more than 20 periods.

Clearly, from a technical point of view, the fact that, near a Hopf bifurcation, the two eigenvalues of the Jacobian matrix are *complex conjugates* with modulus close to one explains the nature of this non-monotonous convergence to the steady-state. Despite of this fact, the impulse response function of output to a sunspot shock do not display the typical hump following transitory shocks characterizing many real economies. Note that all the models considered in Schmitt-Grohé (2000) were unable to account for this feature, and this also includes the Benhabib and Wen (204) model when sunspot shocks alone are considered. Yet, as in this later paper, we expect that the endogenous persistence mechanisms will be strong enough to obtain a hump shaped response of output to transitory demand shocks that are slightly positively correlated (such as preferences shocks, etc.). This is to be verified.

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definition more volatile and would lead to a compression of the IRFs of the other variables. Given the definition of  $u = 1 - l$ , it is clear however that both variables are directly related in the business cycle.

Table 2 - Business Cycle Statistics

Variable (x)	Relative standard deviations with output $\frac{\sigma_x}{\sigma_y}$						
	$c^{ag}$	$c^w$	$c^c$	$i$	$l$	$w$	$WS$
<b>Data</b>	0.56	–	–	3.14	0.93	0.44	0.68
<b>Model</b>	0.66	0.77	0.17	1.88	3.79	3.03	0.24

Variable (x)	Cross correlations with output $Corr(y, x)$						
	$c$	$c^w$	$c^c$	$i$	$l$	$w$	$WS$
<b>Data</b>	0.74			0.90	0.78	0.04	-0.38
<b>Model</b>	0.99	1	0.31	1	1	-0.99	-0.99

	Autocorrelation coefficients on output $\rho_i = corr(y_t, y_{t-i})$					
	$\rho_1$	$\rho_2$	$\rho_3$	$\rho_4$	$\rho_5$	$\rho_6$
<b>Data</b>	0.87	0.68	0.46			
<b>Model</b>	0.89	0.73	0.54	0.33	0.11	-0.10

Finally, we discuss briefly the performance of our model in terms of the standard business cycle statistics emphasized in the Real Business Cycle literature. Table 2 summarizes the main statistics in terms of cross-correlations, relative standard deviations and autocorrelations for the French economy and those implied by the model when submitted to sunspot disturbances of arbitrary size.<sup>29</sup>

Several findings are worth stressing. (to be written)

- The model generates simultaneous *procyclical* movements of aggregate consumption, investment and output. This is worth stressing, as standard business cycle models naturally tend to generate countercyclical movements of consumption or investment in response to demand shocks, unless large increasing returns to scale are introduced.<sup>30</sup>

- Similarly, in contrast to many models, the model does not generate an

<sup>29</sup>Of course, the size of the sunspots must be small enough to ensure that the dynamics remain in the basin of attraction of the steady state.

<sup>30</sup>This issue is discussed in detail in the survey by Benhabib and Farmer (1999) and in Schmitt-Grohe (2000). One solution to mitigate this problem is to combine increasing returns with variable capacity utilization, as in Benhabib and Wen (2004), or endogenous countercyclical markups, as in Dos Santos Ferreira and Dufourt (2006).

excessive smoothness of consumption relative to output.

=> the explanation for these two features stems from the fact that workers are financially constrained and choose to spend all their available income for current consumption. As total income of workers are procyclical (the decrease in individual real wages is more than offset by the increase in employment), so is workers consumption, explaining the procyclicality of consumption and its strong sensitivity to current production

- among the problems : excessive volatility of the employment rate (and of real wages). But we believe this is not a strong limitation. Introducing technological shocks or an intensive margin for labor (effort) would probably strongly mitigate the problem (to be verified).

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